

IVO'S NOTES - 4/24/07 (1)

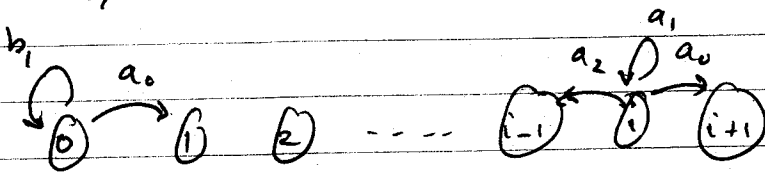
Discrete time M/M/1 or birth-and-death MC.

MC with states $\{0, 1, 2, \dots\}$ and

$$P_{i,i+1} = a_0, \quad P_{i,i-1} = a_2, \quad P_{i,i} = 1 - a_0 - a_2 =: a_1$$

$$P_{0,1} = a_0, \quad P_{0,0} = b_1 (= 1 - a_0)$$

Assume $a_0 < a_2$



$$P = \begin{bmatrix} b_1 & a_0 & 0 & & \\ a_2 & a_1 & a_0 & & \\ & a_2 & a_1 & a_0 & \\ & & & & \ddots \\ 0 & & & & \ddots \end{bmatrix}$$

Example M/M/1 with arrival rate λ and service rate μ

Take $\Delta \geq 0$ small. Then consider # jobs in system at $t = 0, \Delta, 2\Delta, \dots$. In interval Δ there is an arrival with prob $\Delta \lambda =: a_0$ and a departure with prob $\Delta \mu =: a_2$

Let P_i denote steady state prob. What is P_i ?

Note: $\frac{P_j}{P_i}$ = expected # visits to j before first return to i , given that P starts in i .

Define $r_i^{(k)}$ = expected # visits to $i+k$ before first return to i , given that P starts in i .

Because transition prob. are homogeneous and P is skip-free to the right, we have that

$r_i^{(k)}$ does not depend on i ! So: $r_i^{(k)} = r^{(k)}$; Let $r := r^{(1)}$

Further: $P_{i+1} = P_i r^{(1)} = P_i r = \dots = P_0 r^{(i+1)}$, $i = 0, 1, \dots$

What is r ? I.e. how to determine r ? Note: since $P_i \rightarrow 0$, it follows $r < 1$.

Since each excursion reaching $i+k$ first has to pass $i+k-1$ (skip-free!), we have:

$$r^{(k)} = r^{(k-1)} \cdot r^{(1)} = r^{(k-1)} \cdot r = \dots = r^k$$

Condition on the last state visited before visiting $i+1$, we get:

$$r^{(1)} = 1 \cdot a_0 + r \cdot a_1 + r^2 \cdot a_2$$

So, since $r^{(1)} = r$,

$r = a_0 + r a_1 + r^2 a_2$

 (1)

r is characterized as smallest non-negative sol. of (1):

Proof: Let x be smallest non-negative sol. of (1).

How to find x ? Consider iteration

$$X(n+1) = a_0 + X(n) a_1 + X(n)^2 a_2, \quad X(0) = 0$$

Then $X(n) \uparrow$ and $X(n) \leq r$ (by induction).

So $\lim_{n \rightarrow \infty} X(n) =: X(\infty)$ exists, and $X(\infty)$ is less or equal to any nonnegative sol. of (1).

Hence: $X \geq X(\infty)$ and $X \leq r$.

But also $X \geq r$:

Define $r^{(k)}(n)$ = expected # visits to $i+k$ before first return to i over n periods, given that P starts in i .

Thus $r^{(k)}(n) \uparrow r^{(k)}$ as $n \rightarrow \infty$.

Also:

$$r^{(k)}(n) = \sum_{m=0}^n p_{i, i+k-1}^{(m)} r^{(k)}(n-m) \leq \sum_{m=0}^n p_{i, i+k-1}^{(m)} r^{(k)}(n) = r^{(k)}(n) \cdot r$$

↙ ↘
m-steps transition prob. of P .

Hence: $r^{(k)}(n) \leq (r^{(1)})^k$.

Now $r^{(1)}(0) = 0 \leq x(0)$

Suppose $r^{(n)}(n) \leq x(n)$.

$$\begin{aligned}
\text{Then } r^{(n+1)}(n+1) &= a_0 + r^{(n)}(n) a_1 + r^{(n)}(n) a_2 \\
&\leq a_0 + r^{(n)}(n) a_1 + (r^{(n)}(n))^2 a_2 \\
&\leq a_0 + x(n) a_1 + (x(n))^2 a_2 \\
&= x(n+1)
\end{aligned}$$

Since $r^{(n)}(n) \uparrow r^{(n)} = r$, we have $r \leq x$ and thus $r = x$ □

This also gives a mean to determine r :

$$\text{Let } r(0) = 0, \quad r(n+1) = a_0 + r(n) a_1 + r(n)^2 a_2, \quad n=0,1,\dots$$

Then $r(n) \uparrow r$ as $n \rightarrow \infty$.

Remark: we only used homogeneous transition prob. and skip-free to the right.

Hence it also works for

$$P = \begin{pmatrix} b_1 & a_0 & & & \\ b_2 & a_1 & a_0 & & 0 \\ b_3 & a_2 & a_1 & a_0 & \\ \vdots & \vdots & \vdots & \vdots & \vdots \end{pmatrix}, \quad \sum_{i=0}^{\infty} a_i = 1, \quad b_j := 1 - \sum_{i=0}^{j-1} a_i$$

"G/M/1" MC's (i.e. embedded MC of the G/M/1)

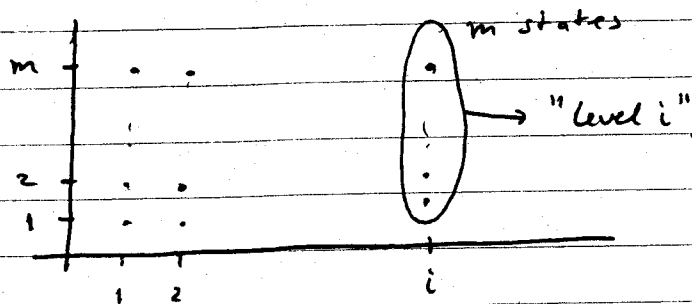
Then r is ~~the~~ minimal non-negative solution of

$$r = a_0 + r a_1 + r^2 a_2 + \dots = \sum_{i=0}^{\infty} r^i a_i$$

Now replace every state by a set of states (of same size)

Quasi birth-and-death process (QBD) or M/M/1-type model

state space:



Irreducible MC with states

$$\left\{ \underbrace{(0,1), \dots, (0,m)}_{\text{level 0}}, \underbrace{(1,1), \dots, (1,m)}_{\text{level 1}}, \dots \right\}$$

\nearrow phase 1 \nearrow phase m

and transition prob. matrix

$$P = \begin{bmatrix} B_{00} & A_0 & & & \\ B_{10} & A_1 & A_0 & & 0 \\ 0 & A_2 & A_1 & A_0 & \\ 0 & 0 & A_2 & A_1 & A_0 \\ & & & \ddots & \ddots \end{bmatrix}$$

repeating block structure
All blocks of size $m \times m$.

Note that $A_2 + A_1 + A_0$ is also MC (i.e. transition prob. matrix)
This MC describe behavior within level (between phases)

Assume P is irreducible and $A_2 + A_1 + A_0$ has 1 communication class.

When is P stable?

Stability (Neut's mean drift condition)

P is positive recurrent if and only if

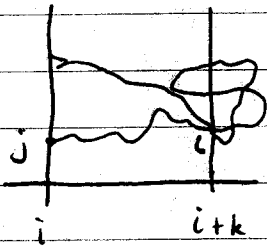
$$\underbrace{\pi A_0 e}_{\text{drift to right}} < \underbrace{\pi A_2 e}_{\text{drift to left}} \quad \text{or: } \underbrace{1 \cdot \pi A_0 e + 0 \cdot \pi A_1 e - 1 \cdot \pi A_2 e}_{\text{mean step size}} < 0 \quad (e = \begin{pmatrix} 1 \\ 1 \\ 1 \end{pmatrix})$$

where π is equilibrium distr. of A_0, A_1, A_2 : $\pi(A_0 + A_1 + A_2) = \pi e = 1$.

Assume P is stable

Let $P(i, j)$ denote steady-state prob. of state (i, j) .

Define: $R_i^{(k)}(j, l) =$ expected # visits to $(i+k, l)$ before first return to level i , given that P starts in (i, j) .



Since P is homogeneous and skip-free to the right: $R_i^{(k)} = R^{(k)}$
 does not depend on i !

and

$$P(i+1, k) = \sum_{j=1}^m P(i, j) \cdot R^{(1)}(j, k)$$

or: $\underbrace{P_{i+1}}_{\text{vector}} = (P(i+1, 1), \dots, P(i+1, m)) = P_i R^{(1)} = P_i R = \dots = P_0 R^i$
 \downarrow
 $R := R^{(1)}$

Note: since $P_i \rightarrow 0$, it follows $\sigma(R) < 1$
 \downarrow
 spectral radius.

How to determine R ?

We have $R^{(k)} = R^{(k-1)} \cdot R^{(1)} = \dots = (R^{(1)})^k = R^k$.

and (condition on last state visited before visiting $(i+1, l)$)

$$R^{(1)} = A_0 + R^{(1)} A_1 + R^{(2)} A_2$$

so $R = A_0 + R A_1 + R^2 A_2$ (2)

Just as before (same proof!): R is minimal non-negative solution of (2).

and R can be determined by iteration:

$R(0) = 0$, $R(n+1) = A_0 + R(n)A_1 + (R(n))^2 A_2$, $n = 0, 1, 2, \dots$
Then $R(n) \uparrow R$ as $n \rightarrow \infty$. (3)

Note: alternative: $R(I - A_1) = A_0 + R^2 A_2$

so $R = (A_0 + R^2 A_2)(I - A_1)^{-1}$
exists = $I + A_1 + A_1^2 + \dots$

So: $R(0) = 0$, $R(n+1) = (A_0 + R^2(n)A_2)(I - A_1)^{-1}$, $n = 0, 1, \dots$

Usually converges faster.

Finally: P_0 follows from boundary eqs.

$$P_0 B_{00} + P_1 B_{10} = P_0$$

$P_1 = P_0 R$
 $\rightarrow P_0 [B_{00} + R B_{10}] = P_0$
equilibrium eqs. of MC embedded on level 0

$R B_{10}(j, l)$ = prob. that P returns to $(0, l)$ from excursion starting in $(0, j)$

Note: It also works for

more complex boundary behavior.

$$P = \begin{bmatrix} B_{00} & A_0 & & & \\ B_{10} & A_1 & A_0 & & 0 \\ B_{20} & A_2 & A_1 & A_0 & \\ \vdots & \vdots & \vdots & \vdots & \vdots \end{bmatrix} \text{ or: } P = \begin{bmatrix} B_{00} & B_{01} & & & \\ B_{10} & B_{11} & A_0 & & 0 \\ B_{20} & A_2 & A_1 & A_0 & \\ \vdots & \vdots & \vdots & \vdots & \vdots \end{bmatrix}$$

- $B_{00}: n \times n$
- $B_{01}: n \times m$
- $B_{i0}: m \times n$
- $A_i: m \times m$

* Then: P is minimal non-negative solution of

$$R = \sum_{i=0}^{\infty} R^i A_i$$

* Drift condition: $\pi A_0 e < \pi (A_1 + 2A_2 + \dots) e = \pi \sum_{i=1}^{\infty} i A_i e$

$$\text{or: } \pi \sum_{i=0}^{\infty} (i-1) A_i e < 0.$$

mean step size.

Now: MC in continuous time (MP)

$$\text{generator } Q = \begin{bmatrix} B_{00} & A_0 & & \\ A_2 & A_1 & A_0 & 0 \\ & A_2 & A_1 & A_0 \\ 0 & & & \ddots \end{bmatrix}$$

(note that $A_2 + A_1 + A_0$ is also generator)

To find $P = (P_0, P_1, \dots)$ consider

$$P := I + \Delta Q, \text{ where } \Delta \geq 0, \text{ sufficiently small such that } P \geq 0.$$

Then P is stochastic matrix and P and Q have the same equilibrium distr:

$$p = pP \iff p = p(I + \Delta Q) \iff 0 = \Delta pQ \iff 0 = pQ$$

This also implies that P is positive recurrent iff Q is positive recu

$$P = \begin{bmatrix} I + \Delta B_{00} & \Delta A_0 & & \\ \Delta A_2 & I + \Delta A_1 & \Delta A_0 & 0 \\ & \Delta A_2 & I + \Delta A_1 & \Delta A_0 \\ 0 & & & \ddots \end{bmatrix}$$

Drift condition: $\pi \Delta A_0 e < \pi \Delta A_2 e$, where $\pi(I + \Delta(A_2 + A_1 + A_0)) = e$
 or: $\pi(A_2 + A_1 + A_0) = e$
 $\pi e = 1$

P_i : $P_i = P_0 R^i$ where R is min. non-neg. sol. of:

$$R = \Delta A_0 + R(I + \Delta A_1) + R^2 \Delta A_2$$

so

$$0 = A_0 + R A_1 + R^2 A_2$$

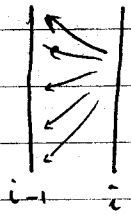
Determine R via (3) or via: $R = (A_0 + R^2 A_2) \underbrace{(-A_1)^{-1}}_{\geq 0!}$

Special cases of M/M/1-type (or QBD) models

(1) $A_2 = V \cdot \alpha = \begin{pmatrix} v_1 \alpha \\ v_2 \alpha \\ \vdots \\ v_m \alpha \end{pmatrix}, \alpha = (\alpha_1, \dots, \alpha_m), \alpha e = 1$
 $V = \begin{pmatrix} v_1 \\ \vdots \\ v_m \end{pmatrix}$

So rows of A_2 are the same upto scaling.

This means: given that P jumps to a lower level, the entrance distribution is always the same, i.e. it does not depend on which state P came from.



What is entrance distr.? $\alpha!$

Note that: v_j is prob. that P jumps to lower level from (i,j)

So $R = A_0 + RA_1 + \underbrace{R^2 A_2}_{RA_0 e \cdot \alpha} \rightarrow R[I - A_1 - A_0 e \alpha] = A_0$

$R = A_0 [I - (A_1 + A_0 e \alpha)]^{-1}$

(in continuous time: $R = -A_0 (A_1 + A_0 e \alpha)^{-1}$)

Alternative derivation:

$P_i = P_{i-1} A_0 + P_i A_1 + P_{i+1} A_2 \stackrel{A_2 = V \alpha}{=} P_{i-1} A_0 + P_i A_1 + P_{i+1} V \alpha$

Balance flow between level i and level $i+1$:

$P_i A_0 e = P_{i+1} A_2 e = P_{i+1} V$

Hence

$P_i = P_{i-1} A_0 + P_i A_1 + P_i A_0 e \alpha$

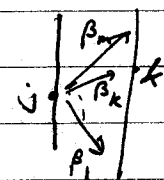
so $P_i = P_{i-1} \underbrace{A_0 [I - A_1 - A_0 e \alpha]^{-1}}_R$

$$(2) A_0 = w \cdot \beta = \begin{pmatrix} w_1 \\ \vdots \\ w_m \end{pmatrix} (\beta_1, \dots, \beta_m), \quad \beta_2 = 1$$

This means: if P jumps to higher level, then the entrance dist. is always the same, i.e., it does not depend on which state P came from.

(note that w_j is the prob. that P jumps to higher level from (i,j))

This implies:



$$R(j,l) = (1-w_j) \cdot 0 + w_j \cdot \underbrace{\sum_{k=1}^m \beta_k S(k,l)}_{\text{only depends on } l} =: a_l$$

where $S(k,l)$ = expected # visits to $(i+1,l)$ before first return to level i , given that P starts in $(i+1,k)$.

So R is of the form:

$$R = w \cdot a \quad \text{for some vector } a = (a_1, \dots, a_m)$$

Alternative derivation: look at iteration scheme

$$R^{(n+1)} = A_0 + R^{(n)} A_1 + R^{(n-1)} A_2, \quad A_2 = w \cdot \beta$$

Then $R^{(1)} = A_0$, so it is of the form $R^{(1)} = w \cdot a_1$
(so $R^{(2)} = (a_1 w) \cdot w \cdot a_1$)

$$\text{and } R^{(2)} = w \left[\beta + a_1 A_1 + (a_1 w) a_1 A_2 \right] = w a_2$$

⋮

$R^{(n)} = w a_n \uparrow R$, so also R is of the form $R = w \cdot a$.

Thus $R^i = (aw)^{i-1} R = \eta^{i-1} R$ where $\eta := aw$

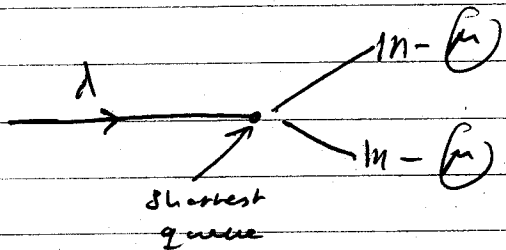
$$\text{so } p_i = p_0 R^i = \eta^{i-1} p_0 R = \eta^{i-1} p_1, \quad i=1,2,\dots$$

and η is unique root on $(0,1)$ of: $\det(A_0 + (A_1 - I)\eta + A_2 \eta^2) = 0$

or determine a via: $a_{n+1} = \beta + a_n A_1 + (a_n w) a_n A_2, n=0,1,\dots, a_0=0$

Examples

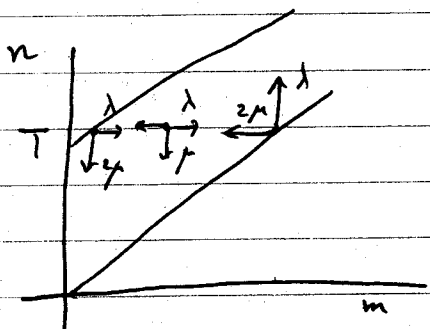
(1) Shortest queue with threshold jockeying



If difference between longest and shortest queue $> T$, then job jumps from longest to shortest queue.

States ~~(m,n)~~ (m,n) where $m =$ length of shortest queue
 $n =$ length of longest queue

(so $m \leq n$)



Levels? Take level $i = \{ \text{states } (n-T, n), (n-T+1, n), \dots, (n, n) \}$

and put states (m,n) with $n \leq T$ into set of boundary states.

Then $A_0 = \begin{pmatrix} 0 \\ \vdots \\ 0 \dots 0 \end{pmatrix}$: only 1 nonzero row!

Hence $P_n = (P(n-T, n), \dots, P(n, n)) = \eta^{n-T} \cdot P_T, \quad n = T, T+1, \dots$

What is η ? Let $V_L = \{ (m,n) \mid m+n = L \}$

Then balancing flow between V_L and V_{L+1} gives

$$P(V_{L+1}) \cdot 2\mu = P(V_L) \cdot \lambda \quad (L \geq T)$$

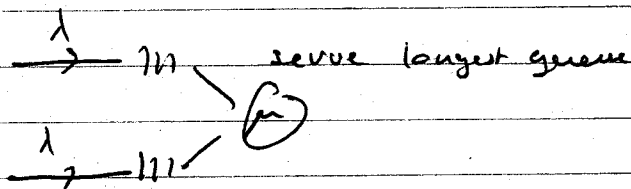
So $P(V_{t+2}) = \rho^2 P(V_t)$

Also $p(m+1, n+1) = \eta \cdot p(m, n)$ and thus

$P(V_{t+2}) = \eta P(V_t)$

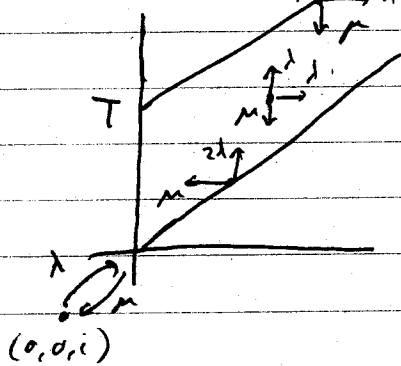
Conclusion : $\eta = \rho^2$.

(2) Longest queue with threshold addition/rejection



If diff between longest queue and shortest queue $> T$,

(threshold rejection) λ (threshold addition)



then :

- (i) add job to shortest queue
- or : (ii) reject job in longest queue

States (m, n) ,

$m =$ length shortest queue

$n =$ length longest queue

$(0,0,i) =$ queues are empty, server idle.

Levels? Take level $m = \{ (m, m), (m, m+1), \dots, (m, m+T) \}$

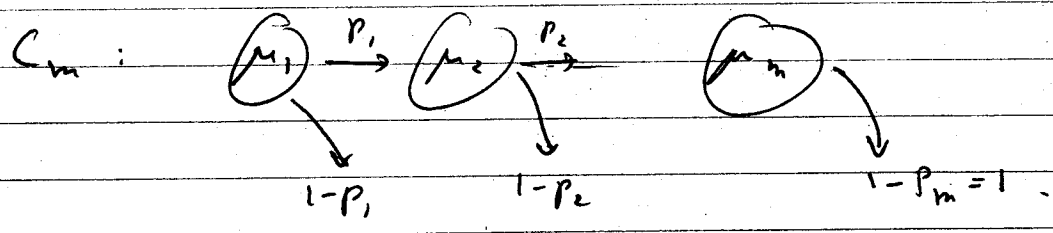
Then $A_2 = \begin{pmatrix} 0 & \mu & 0 & 0 \\ & & 0 & \\ & & & \\ & & & \end{pmatrix}$; only 1 non-zero row!

Hence

$R = -A_0 (A_1 + A_0 e \alpha)^{-1}$

where $v = \begin{pmatrix} \mu \\ 0 \\ \vdots \\ 0 \end{pmatrix}$, $\alpha = (0, 1, 0, \dots)$

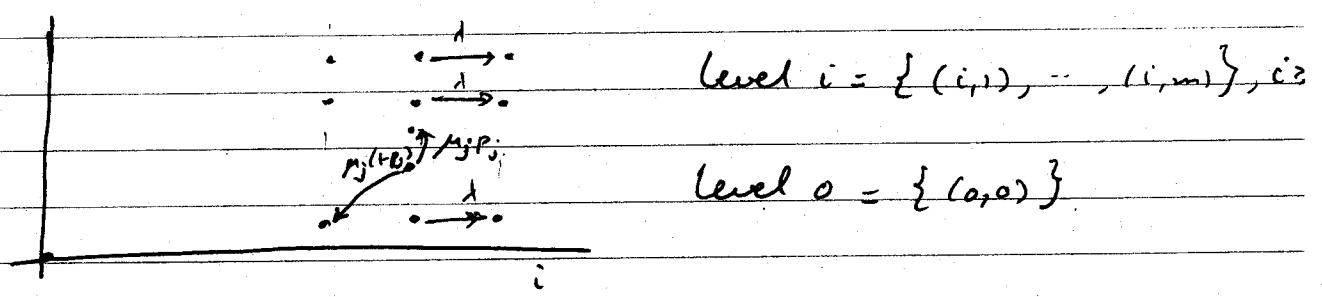
(3) $M/C_m/1$ queue



Class of distributions C_m is dense:

every (continuous) distr. on $(0, \infty)$ can be approximated arbitrarily close by C_m -distr.

States (i, j) : $i = \#$ jobs in system
 $j =$ phase of job in service, $j = 1, \dots, m$
 0 : idle (empty) state.



So: $A_0 = \lambda I$, $A_1 = \begin{pmatrix} -(\lambda + \mu_1) & \mu_1 p_1 & & 0 \\ & \ddots & \ddots & \\ & & \mu_m p_m & \\ 0 & & & -(\lambda + \mu_m) \end{pmatrix}$

$A_2 = \begin{pmatrix} \mu_1(1-p_1) & & & \\ & \ddots & & \\ & & 0 & \\ & & & \mu_m(1-p_m) \end{pmatrix} = V \cdot \alpha$ where $V = \begin{pmatrix} \mu_1(1-p_1) \\ \vdots \\ \mu_m(1-p_m) \end{pmatrix}$
 $\alpha = (1, 0, \dots, 0)$

Hence: $R = -A_0(A_1 + A_2 \alpha)^{-1} \cdot \alpha$